

STATS 207: Time Series Analysis

Autumn 2020

Lecture 2: Sample ACF and Basic Theoretical Constructs

Dr. Alon Kipnis

Slides credit: David Donoho, Dominik Rothenhäusler

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WHITE NOISE

AUTOVARIANCE AND AUTOCORRELATION

STATIONARITY

ESTIMATING AUTO- AND CROSSCOVARIANCE

Genera Info

- Home assignment 1 will be posted on Monday 9/21/2020. It is Due two weeks later (Monday 10/5/2020).
- We will announce on Canvas and provide the link as-soon-as the assignment is posted.
- Lecture 1 slides and recording are available on Canvas.
- Syllabus with estimated week numbers for each topic is on Canvas.

Models for Time Series Data

Name	Example
White noise	$w_t \sim \mathcal{N}(0, \sigma^2)$
Moving Average	$x_t = (w_{t-1} + w_t + w_{t+1})/3$
Autoregression	$x_t = x_{t-1} - 0.9x_{t-2} + w_t$
Random Walk	$x_t = x_{t-1} + w_t$
Sinusoid in noise	$x_t = 2 \cos(2\pi t/50 + 0.6\pi) + w_t$

Recall lowercase notation: For each $t = 0, 1, \dots$, x_t , is a random variables.

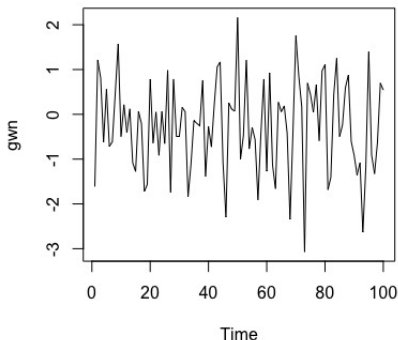
White Noise

3 Flavors of White Noise

A **Change** to the definitions in lecture 1:

- **Definition:** (w_t) is *white noise* if
 1. $\mathbb{E}[w_t] = 0$ for all $t = 1, 2, \dots$ (zero mean)
 2. $\mathbb{E}[w_t^2] = \sigma_w^2$ for all $t = 1, 2, \dots$ (finite and identical variance)
 3. $\mathbb{E}[w_s w_t] = 0$, for all $t \neq s$ (pairwise uncorrelated).
- **Definition:** (w_t) is *iid noise* if it is white noise and if w_1, w_2, \dots are independent and identically distributed.
- **Definition:** (w_t) is *white Gaussian noise* if it is iid noise and $w_t \sim N(0, \sigma_w^2)$.

Gaussian White Noise



- White noise has **no varying structure over time**.
- For most real time series data white noise is not a good model!
- Standard working pipeline:
 1. Transform data
 2. Fit model to data
 3. Check whether residuals are white noise

How to check whether white noise is a good model for data?

Autocorrelation Function

Data: x_1, x_2, \dots, x_n .

- **Definition:** *Sample ACF:*

$$\hat{\rho}(h) \equiv \frac{\sum_{t=1}^{n-h} (x_t - \bar{x})(x_{t+h} - \bar{x})}{\sum_{t=1}^n (x_t - \bar{x})^2}, \quad h = 1, 2, \dots, n-1.$$

- **For white noise** (from the CLT):

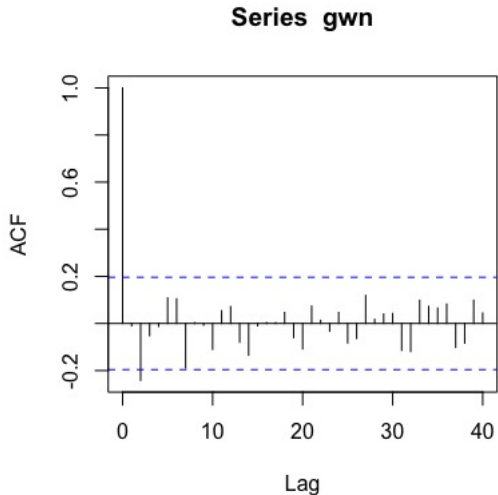
$$\hat{\rho}(h) \xrightarrow{D} \mathcal{N}(0, 1),$$

hence,

$$\Pr(|\hat{\rho}(h)| > 1.96/\sqrt{n}) \approx \Pr(|\mathcal{N}(0, 1)| > 1.96) = 0.05.$$

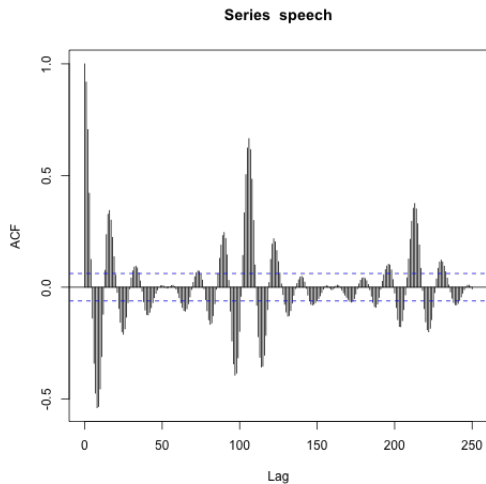
Correlogram: Plot the sample ACF for $h = 1, \dots, k$ to check if white noise is a good model for data (R function `acf`).

Example of Sample ACF of White Noise



ACF of Speech Data

Example 1.27 in [\[Shumway & Stoffer\]](#)



Not a white noise!

Autocovariance and Autocorrelation

Joint random variables

- **Definition:** Joint CDF

$$F_{t_1, \dots, t_n}(c_1, \dots, c_n) = \Pr(x_{t_1} \leq c_1, \dots, x_{t_n} \leq c_n)$$

- **Example:** white Gaussian noise $x_j \stackrel{iid}{\sim} \mathcal{N}(0, 1)$,

$$F_{t_1, \dots, t_n}(c_1, \dots, c_n) = \prod_{i=1}^n \phi(c_i), \quad \phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-t^2/2} dt$$

- **Definition:** *Marginal* distribution function

$$F_t(x) = \Pr(x_t \leq x), \quad x \in \mathbb{R}$$

- **Definition:** Marginal *density* function

$$f_t(x) = \frac{d}{dx} F_t(x).$$

Mean Function

- **Definition:** The *mean* function

$$\mu_{xt} \equiv \mathbb{E}[x_t] = \int_{-\infty}^{\infty} x_t f_t(x) dx$$

- **Example:** (w_t) white noise,

$$\mu_{wt} = 0, \quad \forall t.$$

- Properties:

- Mean is linear:

$$y_t = a \cdot x_t + b \cdot z_t \Rightarrow \mu_{yt} = a\mu_{xt} + b\mu_{zt}$$

- **Warning:** μ_{xt} is **not** the long-run average! Mean at different t can in principle be different.

Example Mean Functions

- Moving Average: $x_t = (w_{t-1} + w_t + w_{t+1})/3$,

$$\begin{aligned}\mu_{xt} &= \mathbb{E}[x_t] \\ &= \mathbb{E}[(w_{t-1} + w_t + w_{t+1})/3] \\ &= (\mathbb{E}[w_{t-1}] + \mathbb{E}[w_t] + \mathbb{E}[w_{t+1}])/3 = 0\end{aligned}$$

- Random walk with drift: $x_t = \delta \cdot t + \sum_{u=1}^t w_u$,

$$\mu_{xt} = \mathbb{E}[x_t] = \mathbb{E}\left[\delta \cdot t + \sum_{u=1}^t w_u\right] = \mathbb{E}[\delta \cdot t] + \mathbb{E}\left[\sum_{u=1}^t w_u\right] = \delta \cdot t.$$

- Signal plus noise: $x_t = 2 \cos(2\pi t/50 + 0.6\pi) + w_t$,

$$\mu_{xt} = 2 \cos(2\pi t/50 + 0.6\pi)$$

Autocovariance Function

- Definition:

$$\gamma_x(s, t) \equiv \text{Cov}(x_s, x_t) = \mathbb{E}[(x_s - \mu_{x_s})(x_t - \mu_{x_t})]$$

- Example: $w_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma_w^2)$,

$$\gamma_w(s, t) = \begin{cases} \sigma^2 & t = s, \\ 0 & t \neq s. \end{cases}$$

- Properties:

- $\gamma_x(t, t) = \text{Var}(x_t) = \mathbb{E}[(x_t - \mu_{x_t})^2]$
- Autocovariance is bilinear:

$$y_t = ax_t + bz_t \Rightarrow \gamma_y(t, t) = a^2\gamma_x(t, t) + b^2\gamma_z(t, t)$$

if every x_t and z_t are uncorrelated.

- Cauchy-Schwartz inequality:

$$\gamma_x(s, t) \leq \sqrt{\gamma_x(t, t)\gamma_x(s, s)}$$

- **Warning:** $\gamma_x(t, t)$ is **not** the long-run variance! Variance at different t can in principle be different.

Autocorrelation Function

- Definition:

$$\rho_x(s, t) \equiv \text{corr}(x_t, x_s) = \frac{\text{Cov}(x_s, x_t)}{\sqrt{\text{Var}(x_t)\text{Var}(x_s)}} = \frac{\gamma_x(s, t)}{\sqrt{\gamma_x(t, t)\gamma_x(s, s)}}.$$

- Example: $w_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$,

$$\gamma_w(s, t) = \begin{cases} 1 & t = s, \\ 0 & t \neq s. \end{cases}$$

- Properties:

- $\rho_x(t, t) = 1$
- $-1 \leq \rho_x(s, t) \leq 1$
- $|\rho_x(s, t)| = 1$ implies perfect linear relationship:

$$x_t = a \cdot x_s + b, \quad \text{for some scalars } a, b$$

Example Autocovariance Functions

- Moving Average: $x_t = (w_{t-1} + w_t + w_{t+1})/3$,

$$\gamma_x(s, t) = \sigma_w^2 \begin{cases} 3/9 & t = s, \\ 2/9 & |t - s| = 1, \\ 1/9 & |t - s| = 2, \\ 0 & |t - s| > 2. \end{cases}$$

Note: averaging/filtering introduces correlation.

- Random Walk with drift: $x_t = \delta \cdot t + \sum_{u=1}^t w_u$,

$$\gamma_x(s, t) = \min(s, t) \sigma_w^2.$$

- Signal plus noise: $x_t = 2 \cos(2\pi t/50 + 0.6\pi) + w_t$,

$$\gamma_x(s, t) = \begin{cases} \sigma_w^2 & t = s, \\ 0 & t \neq s. \end{cases}$$

Stationarity

Stationarity

- *Strict stationarity*:

- **Definition**: for all t_1, \dots, t_n and h ,

$$\{x_{t_1}, \dots, x_{t_n}\} =_D \{x_{t_1+h}, \dots, x_{t_n+h}\},$$

namely

$$F_{t_1, \dots, t_n}(c_1, \dots, c_n) = F_{t_1+h, \dots, t_n+h}(c_1, \dots, c_n).$$

- *stationarity* (aka *weak stationarity*, *mean-variance stationarity*):

- **Definition**: for all t, s ,

$$\mu_{xt} = \mu_{xs} \quad \gamma_x(s, t) = \gamma_x(0, |t - s|).$$

- **Strict stationarity implies weak stationarity.**
- Notation (abuse of):

$$h = t - s, \quad \gamma(h) \equiv \gamma(0, |t - s|)$$

Examples of stationarity (or not)

- **YES**: Moving average $x_t = (w_{t-1} + w_t + w_{t+1})/3$,

$$\mu_{xt} = 0, \quad \gamma_x(s, t) = \sigma_w^2 \begin{cases} 3/9 & t = s, \\ 2/9 & |t - s| = 1, \\ 1/9 & |t - s| = 2, \\ 0 & |t - s| > 2. \end{cases}$$

- **NO**: Random walk with drift $x_t = \delta \cdot t + \sum_{u=1}^t w_u$,

$$\mu_{xt} = \delta \cdot t, \quad \gamma_x(s, t) = \min(s, t) \sigma_w^2.$$

- **NO**: Signal plus noise $x_t = 2 \cos(2\pi t/50 + 0.6\pi) + w_t$,

$$\mu_{xt} = 2 \cos(2\pi t/50 + 0.6\pi), \quad \gamma_x(h) = \begin{cases} \sigma_w^2 & h = 0, \\ 0 & h \neq 0. \end{cases}$$

Joint Stationarity & Cross-Covariance

- **Definition:** *Cross-covariance function* :

$$\gamma_{xy}(s, t) \equiv \text{Cov}(x_s, y_t) = \mathbb{E}[(x_s - \mu_{x_s})(y_t - \mu_{y_t})].$$

Note: $\gamma_{xx}(s, t) = \gamma_x(s, t)$.

- **Definition:** x and y are *jointly stationary* if each is stationary, and

$$\gamma_{xy}(s, t) = \gamma_{xy}(0, s - t)$$

Notation: $\gamma_{xy}(h) = \gamma_{xy}(t + h, t)$.

- **Example:** $x_t = w_t + w_{t-1}$, $y_t = w_t - w_{t-1}$, (w_t) white noise.

$$\gamma_{xy}(s, t) = \sigma_w^2 \begin{cases} 0 & t = s, \\ 1 & s - t = 1, \\ -1 & s - t = -1, \\ 0 & |t - s| \geq 2. \end{cases}$$

Example: Cross-Covariance of Delayed Series

- (x_t) is stationary and zero mean. (y_t) is delayed, noisy version of (x_t) :

$$y_t = Ax_{t-l} + w_t,$$

where (w_t) is white noise independent of (x_t) . Then

$$\gamma_{yx}(h) = A \cdot \gamma_x(h-l) + \gamma_{wx}(h) = A \cdot \gamma_x(h-l)$$

- $\gamma_{yx}(h)$ peaks at $h = l$, takes value $A \cdot \gamma_x(0)$.
- The best linear prediction of y_{t+l} given x_t is

$$\hat{y}_{t+l} = \frac{\gamma_{yx}(l)}{\gamma_x(0)} x_t = Ax_t.$$

- **Definition:** (x_t) is a *Gaussian process* if for any $\{t_1, \dots, t_n\} \subset \mathbb{N}$, $(x_{t_1}, \dots, x_{t_n})$ has a multivariate normal distribution.

Corollary

Every stationary Gaussian process is strictly stationary.

- **Example:** Gaussian random walk $x_t = x_{t-1} + w_t$
 - Is it Gaussian? **Yes**
 - Is it stationary? **No**
 - How about $y_t = x_{t+h} - x_t$, $h \in \mathbb{N}$? **Yes**

- **Definition:** (stationary) *linear process* is the output of a linear time-invariant system

$$x_t = \mu + \sum_{h=-\infty}^{\infty} \psi_h w_{t-h}, \quad \sum_{h=-\infty}^{\infty} |\psi_h| < \infty.$$

Theorem (Cramèr)

Every stationary Gaussian process that is indeterministic is a linear process with Gaussian (w_t).

Examples of Linear Processes

- **Example:** Stationary linear processes:
 - White noise
 - Moving average
 - Non-explosive autoregressive
- **Example:** Non-Gaussian linear process: (U_t) is a sequence of Bernoulli trials (coin tosses $H \rightarrow -1, T \rightarrow 1$).

$$x_t = \mu + \sum_{h=-\infty}^{\infty} \psi_h U_{t-h}$$

Contrasting Example

(w_t) is white Gaussian noise.

$$y_t = \begin{cases} w_t & t \text{ is even} \\ \frac{w_{t-1}^2 - \sigma_w^2}{\sqrt{2}} & t \text{ is odd.} \end{cases}$$

Process (y_t) is:

- Pairwise uncorrelated **Yes**.
- Pairwise independent **No**.
- Strictly stationary **No**.
- Weakly stationary **Yes**.
- Gaussian **No**.

Estimating Auto- and Crosscovariance

Sample Auto/Cross Covariance/Correlation

- **Definition:** *sample mean*

$$\bar{x} \equiv \frac{1}{n} \sum_{t=1}^n x_t$$

- **Definition:** *sample autocovariance function*

$$\hat{\gamma}_x(h) \equiv \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x}).$$

- **Definition:** *sample cross-covariance function*

$$\hat{\gamma}_{xy}(h) \equiv \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(y_t - \bar{y}).$$

- **Definition:** *sample cross-correlation function*

$$\hat{\rho}_{xy}(h) \equiv \frac{\hat{\gamma}_{xy}(h)}{\sqrt{\hat{\gamma}_x(0)\hat{\gamma}_y(0)}}.$$

- **Warning:** mind the difference between *ensemble* averaging (many realizations) and *time* averaging (one realization).

Confidence Limits for Autocovariance

For white noise (w_t):

- Sample autocorrelation:

$$\text{SE} [\hat{\rho}_w(h)] \approx \frac{1}{\sqrt{n}}.$$

- Sample cross-correlation:

$$\text{SE} [\hat{\rho}_{wx}(h)] \approx \frac{1}{\sqrt{n}},$$

where (x_t) is independent of (w_t).

(approximations assume $h \ll n$)

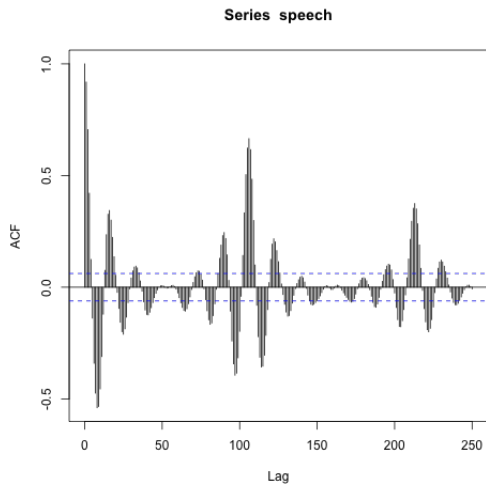
- From the CLT (Theorem A.7 and Property P1.1 in [\[Shumway & Stoffer\]](#)): for $h > 0$,

$$\Pr (|\hat{\rho}_w(h)| > 1.96/\sqrt{n}) \approx \Pr (|\mathcal{N}(0, 1)| > 1.96) = 0.05.$$

$$\Pr (|\hat{\rho}_{wx}(h)| > 1.96/\sqrt{n}) \approx \Pr (|\mathcal{N}(0, 1)| > 1.96) = 0.05.$$

ACF of Speech Data

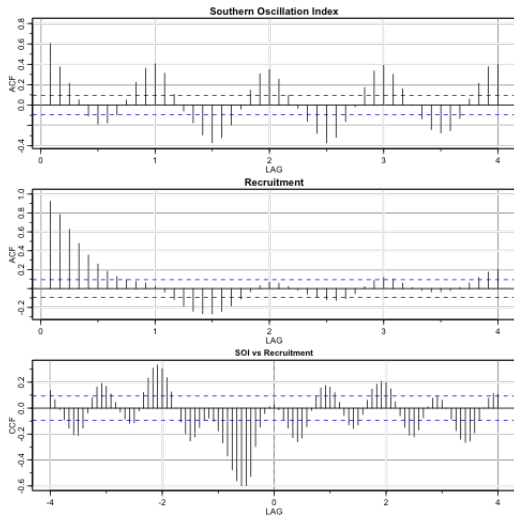
Example 1.27 in [\[Shumway & Stoffer\]](#)



Not a white noise!

ACF/CCF of SOI/Recruitment Data

Example 1.28 in [\[Shumway & Stoffer\]](#)



Multiple Testing Warning

- If $|\hat{\rho}_x(h)| > 1.96n^{-1/2}$ for some $h > 0$, can we determine that (x_t) is not a white noise?

